

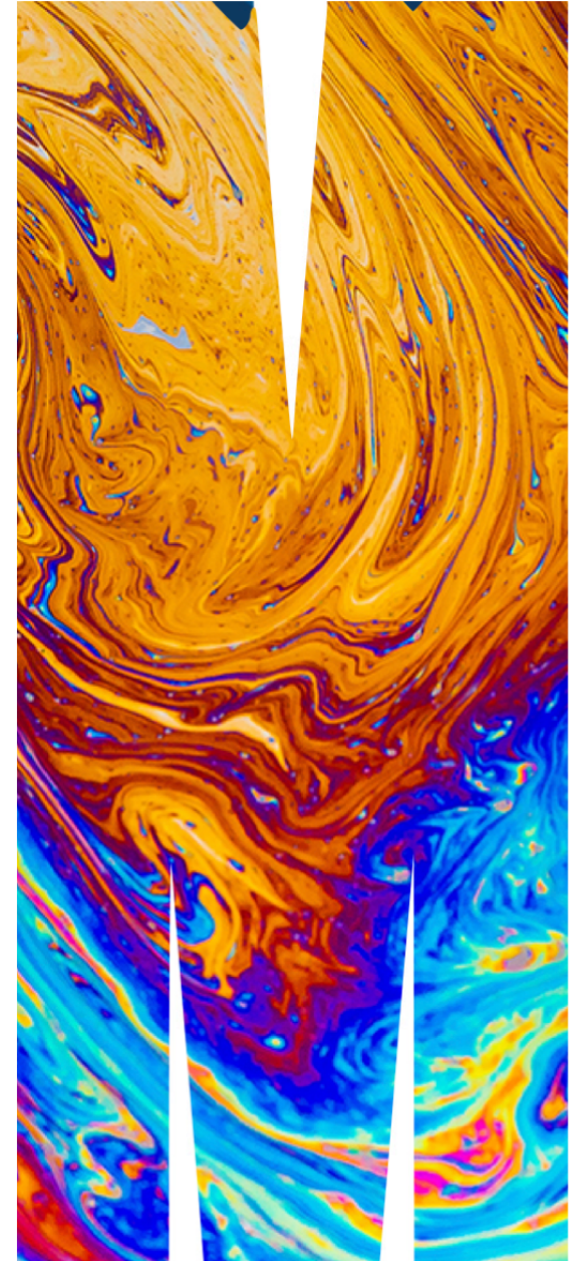
ETC5521: Exploratory Data Analysis

Sculpting data using models, checking assumptions, co-dependency and performing diagnostics

Lecturer: *Di Cook*

✉ ETC5521.Clayton-x@monash.edu

📅 Week 11 - Session 2



Revisiting outliers

- We defined outliers in week 4 as "observations that are significantly different from the majority" when studying univariate variables.
- There is actually no hard and fast definition.



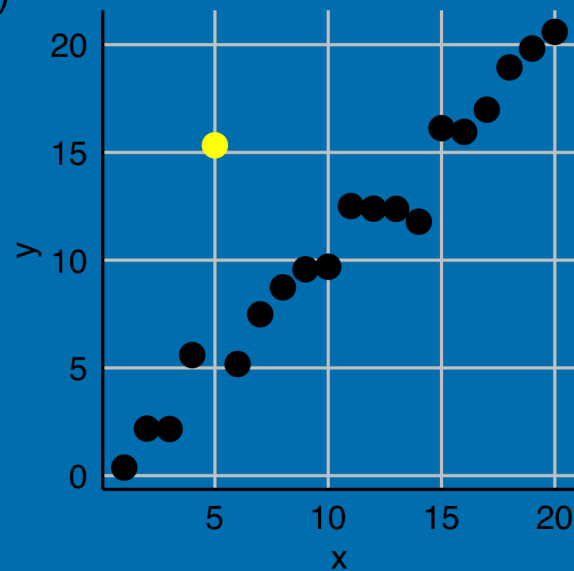
We can also define an outlier as a data point that emanates from a different model than do the rest of the data.

- Notice that this makes this definition *dependent on the model* in question.

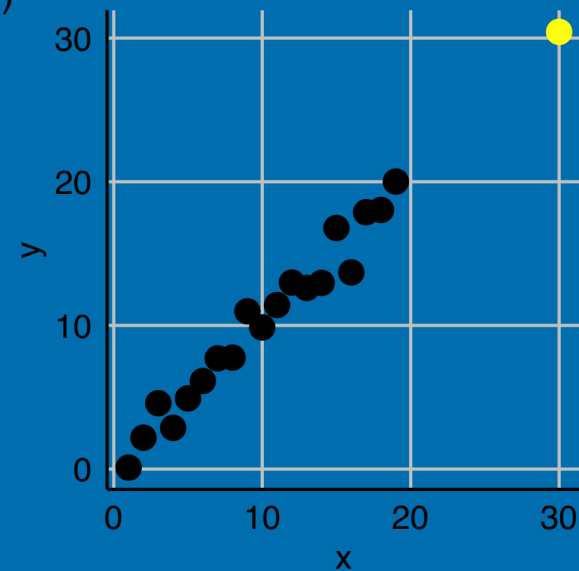
Pop Quiz

Would you consider the yellow points below as outliers?

(A)

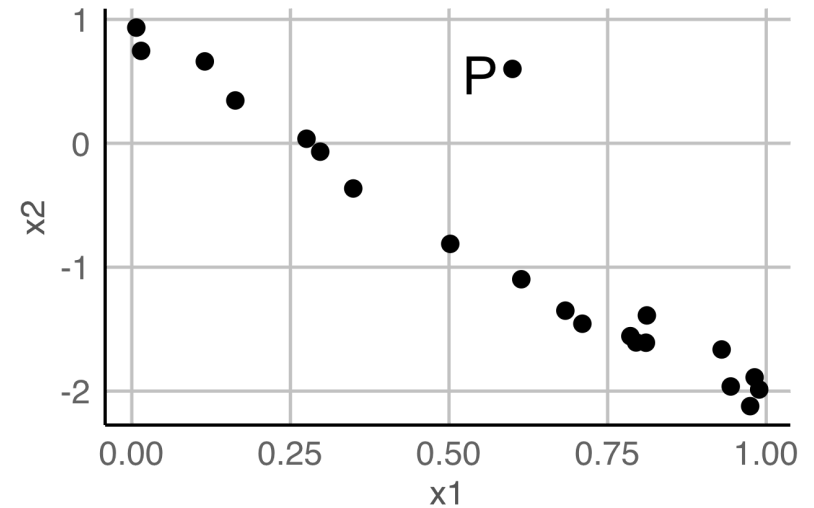


(B)



Outlying values

- As with simple linear regression the fitted model should not be used to predict Y values for x combinations that are well away from the set of observed x_1 values.
- This is not always easy to detect!
- Here, a point labelled P has x_1 and x_2 coordinates well within their respective ranges but P is not close to the observed sample values in 2-dimensional space.
- In higher dimensions this type of behaviour is even harder to detect but we need to be on guard against extrapolating to extreme values.



Leverage

- The matrix $\mathbf{H} = \mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top$ is referred to as the **hat matrix**.
- The i -th diagonal element of \mathbf{H} , h_{ii} , is called the **leverage** of the i -th observation.
- Leverages are always between zero and one,

$$0 \leq h_{ii} \leq 1.$$

- Notice that leverages are not dependent on the response!
- Points with high leverage can exert a lot of influence on the parameter estimates

Leverage

On the data from the previous slide:

```
example_data
```

```
## # A tibble: 21 × 3
##       id      x1      x2
##   <int> <dbl>  <dbl>
## 1     1     1 0.982 -1.89
## 2     2     2 0.297 -0.0679
## 3     3     3 0.115  0.661
## 4     4     4 0.163  0.345
## 5     5     5 0.944 -1.96
## 6     6     6 0.795 -1.61
## 7     7     7 0.975 -2.12
## 8     8     8 0.349 -0.365
## 9     9     9 0.502 -0.812
## 10    10    10 0.810 -1.61
## # i 11 more rows
```

Leverage

```
x <- as.matrix(example_data[2:3])
hat_matrix <- x %*% solve(t(x) %*% x) %*% t(x)
example_data %>%
  mutate(leverage = diag(hat_matrix)) %>%
  print(n = 21)
```

```
## # A tibble: 21 × 4
```

```
##       id      x1      x2 leverage
##   <int>  <dbl>  <dbl>    <dbl>
## 1      1  0.982 -1.89     0.105
## 2      2  0.297 -0.0679  0.0422
## 3      3  0.115  0.661    0.118
## 4      4  0.163  0.345    0.0656
## 5      5  0.944 -1.96     0.106
## 6      6  0.795 -1.61     0.0724
## 7      7  0.975 -2.12     0.123
## 8      8  0.349 -0.365    0.0230
## 9      9  0.502 -0.812    0.0275
```

Studentized residuals

- In order to obtain residuals with equal variance, many texts recommend using the **studentised residuals**

$$R_i^* = \frac{R_i}{\hat{\sigma} \sqrt{1 - h_{ii}}}$$

for diagnostic checks.

Cook's distance

- Cook's distance, D , is another measure of influence:

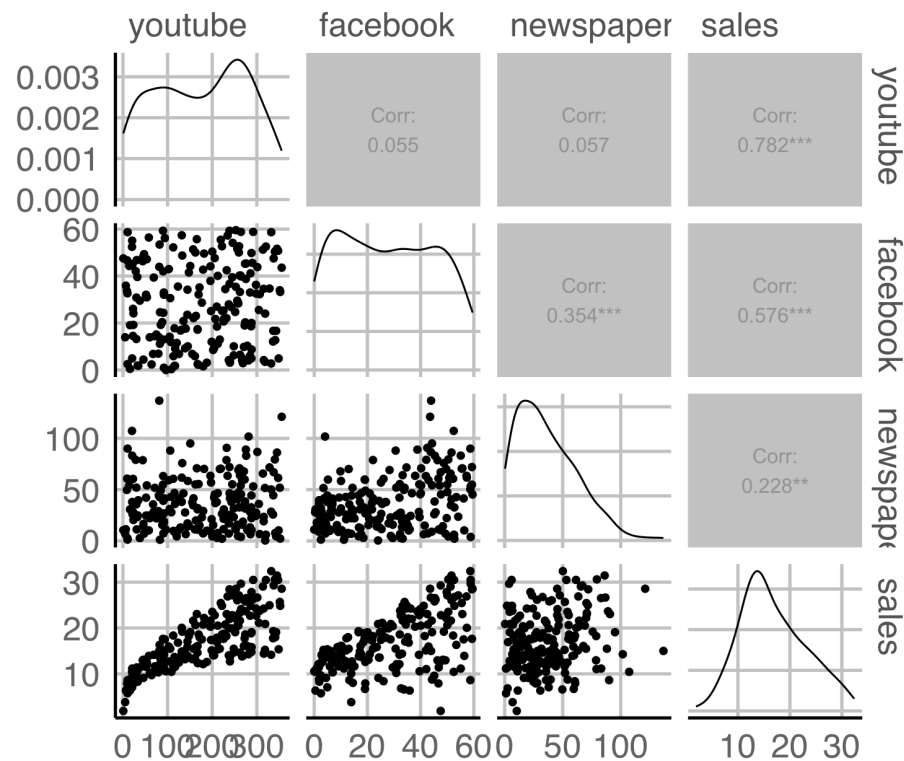
$$\begin{aligned} D_i &= \frac{(\hat{\boldsymbol{\beta}} - \hat{\boldsymbol{\beta}}_{[-i]})^\top \text{Var}(\hat{\boldsymbol{\beta}})^{-1} (\hat{\boldsymbol{\beta}} - \hat{\boldsymbol{\beta}}_{[-i]})}{p} \\ &= \frac{R_i^2 h_{ii}}{(1 - h_{ii})^2 p \sigma^2}, \end{aligned}$$

where p is the number of elements in $\boldsymbol{\beta}$, $\hat{\boldsymbol{\beta}}_{[-i]}$ and $\hat{Y}_{j[-i]}$ are least squares estimates and the fitted value obtained by fitting the model ignoring the i -th data point (\mathbf{x}_i, Y_i) , respectively.

Case study 2 Social media marketing

Data collected from advertising experiment to study the impact of three advertising medias (youtube, facebook and newspaper) on sales.

 data R



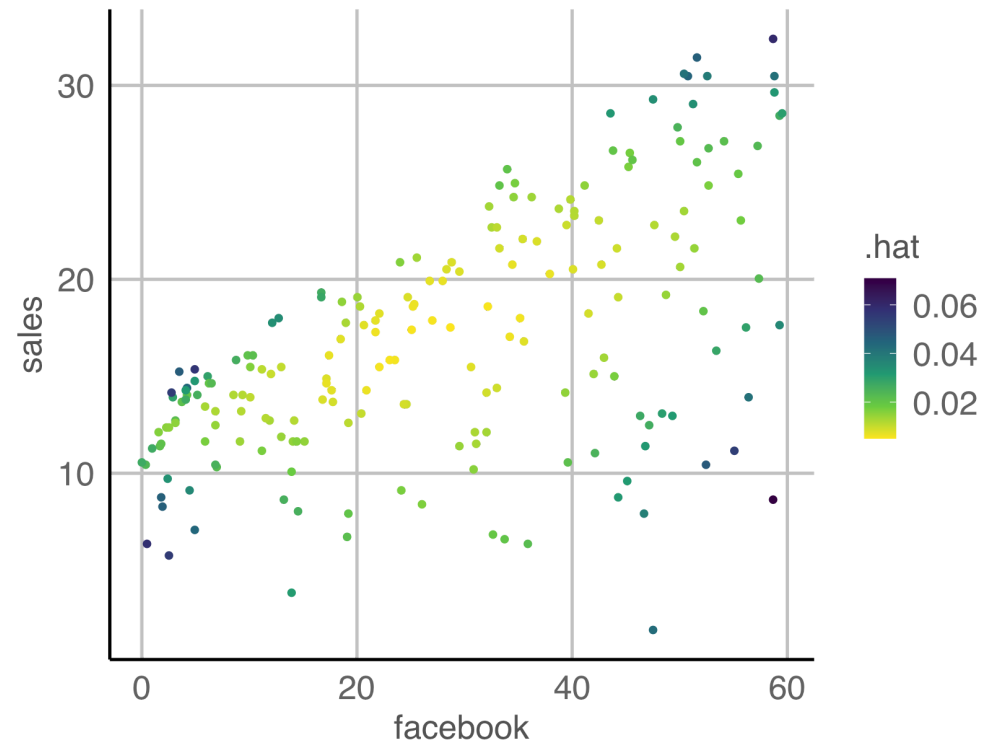
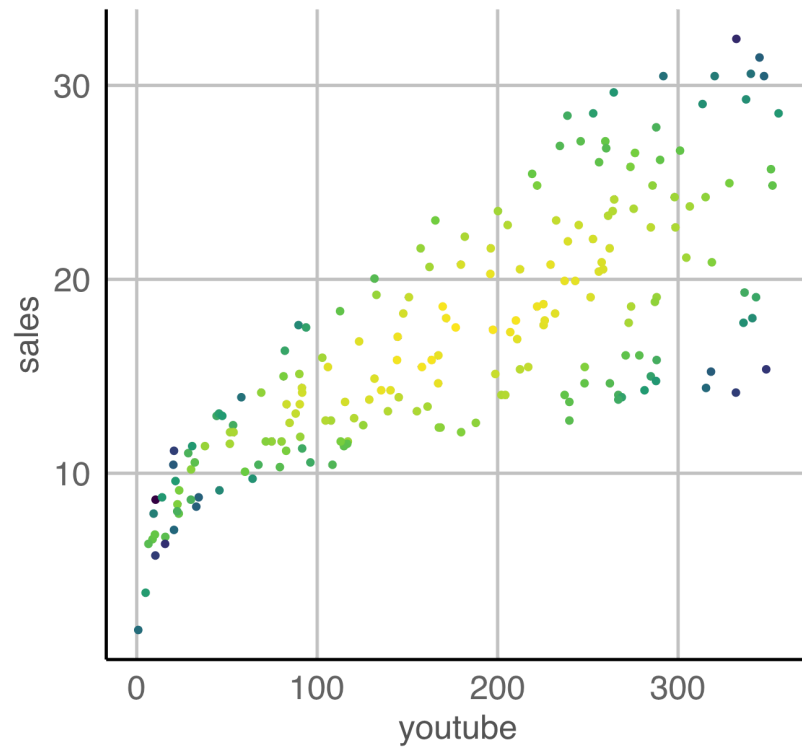
Extracting values from models in R

- The leverage value, studentised residual and Cook's distance can be easily extracted from a model object using `broom::augment`.
 - `.hat` is the leverage value
 - `.std.resid` is the studentised residual
 - `.cooks` is the Cook's distance

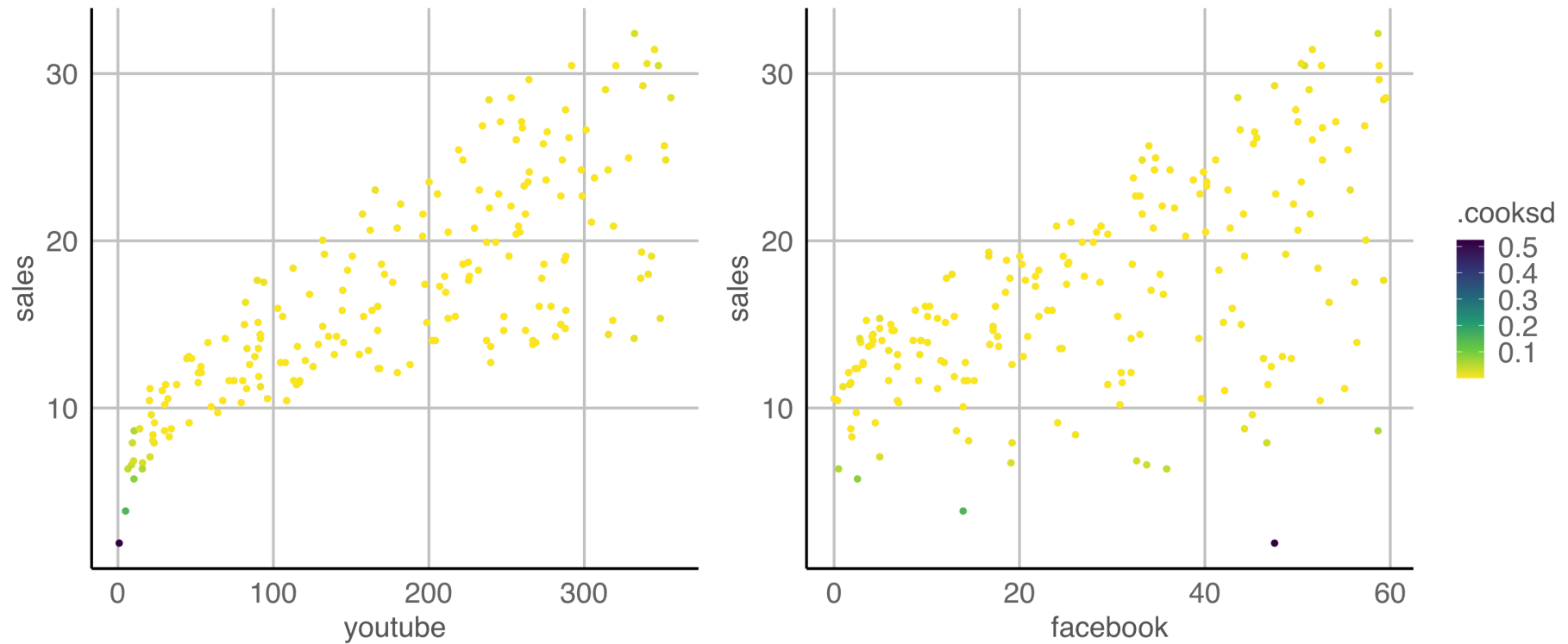
```
fit <- lm(sales ~ youtube * facebook, data = marketing)
(out <- broom::augment(fit))
```

```
## # A tibble: 200 × 9
##   sales youtube facebook .fitted .resid   .hat .sigma   .cooks .std.resid
##   <dbl>   <dbl>   <dbl>   <dbl>   <dbl>   <dbl> <dbl>   <dbl>   <dbl>
## 1 26.5    276.    45.4    26.0    0.496  0.0174  1.13  0.000864    0.442
## 2 12.5     53.4    47.2    12.8   -0.281  0.0264  1.13  0.000431   -0.252
## 3 11.2     20.6    55.1    11.1    0.0465  0.0543  1.14  0.0000256    0.0423
## 4 22.2    182.    49.6    21.2    1.04   0.0124  1.13  0.00268     0.923
## 5 15.5    217.    13.0    15.2    0.316  0.0104  1.13  0.000207    0.280
```

Examining the leverage values



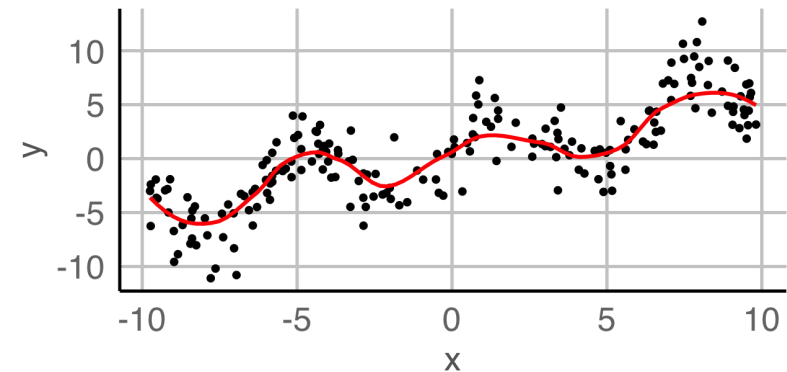
Examining the Cook's distance



Non-parametric regression

LOESS

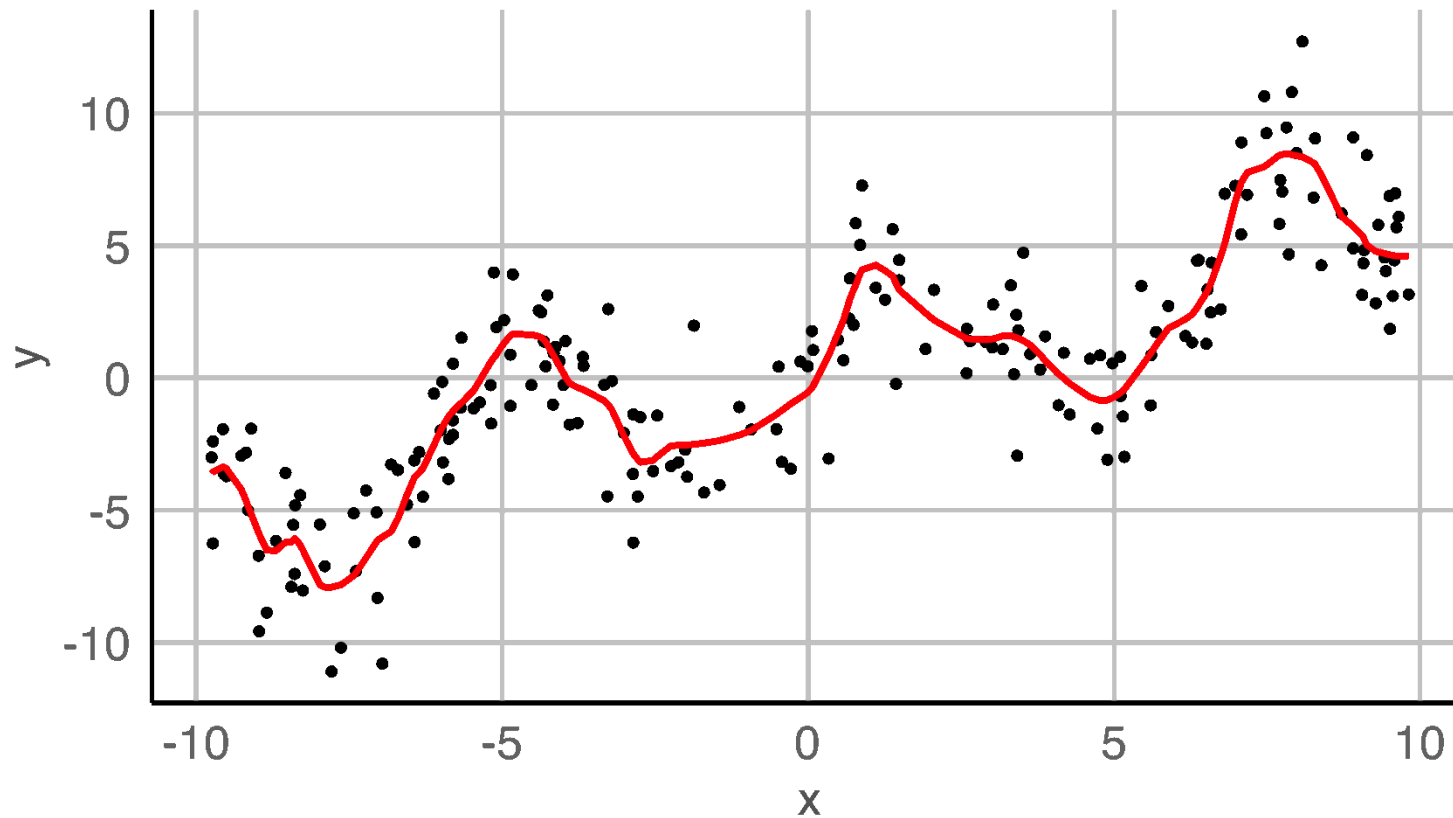
- LOESS (LOcal regrESSion) and LOWESS (LOcally WEighted Scatterplot Smoothing) are **non-parametric regression** methods (LOESS is a generalisation of LOWESS)
- **LOESS fits a low order polynomial to a subset of neighbouring data** and can be fitted using `loess` function in `R`
- a user specified "bandwidth" or "smoothing parameter" α determines how much of the data is used to fit each local polynomial.



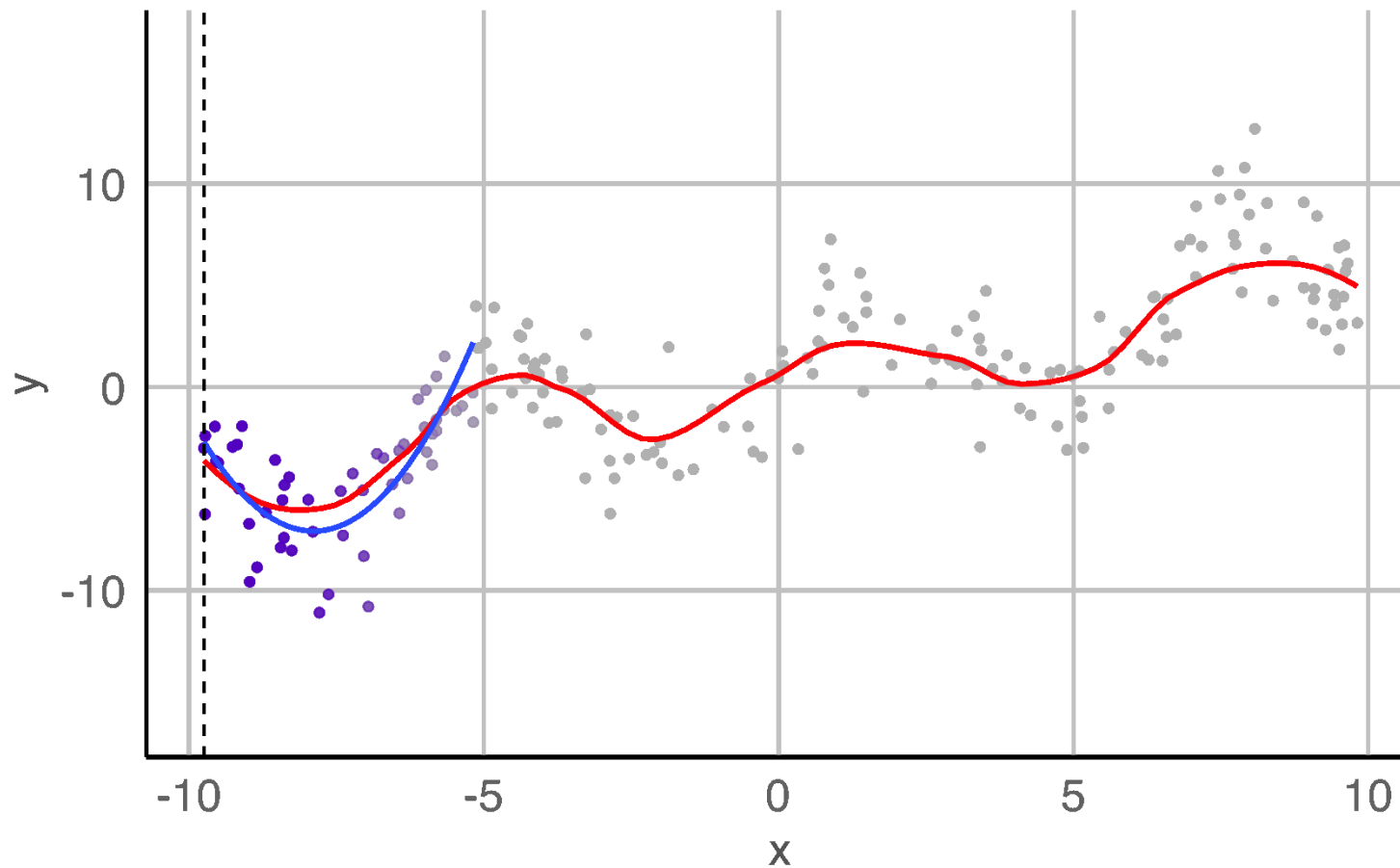
- $\alpha \in \left(\frac{\lambda+1}{n}, 1 \right)$ (default `span=0.75`) where λ is the degree of the local polynomial (default `degree=2`) and n is the number of observations.
- Large α produce a smoother fit.
- Small α overfits the data with the fitted regression capturing the random error in the data.

How **span** changes the loess fit

span = 0.1



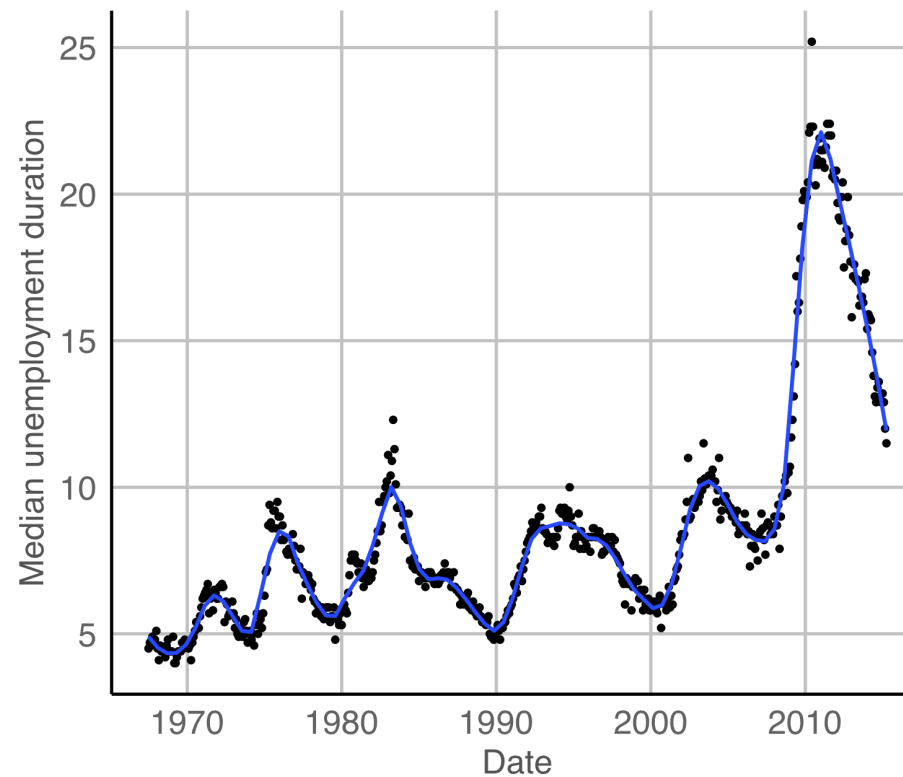
How **loess** works



Case study 3 US economic time series

This dataset was produced from US economic time series data available from <http://research.stlouisfed.org/fred2>.

 data R



How to fit LOESS curves in R?

Model fitting

The model can be fitted using the `loess` function where

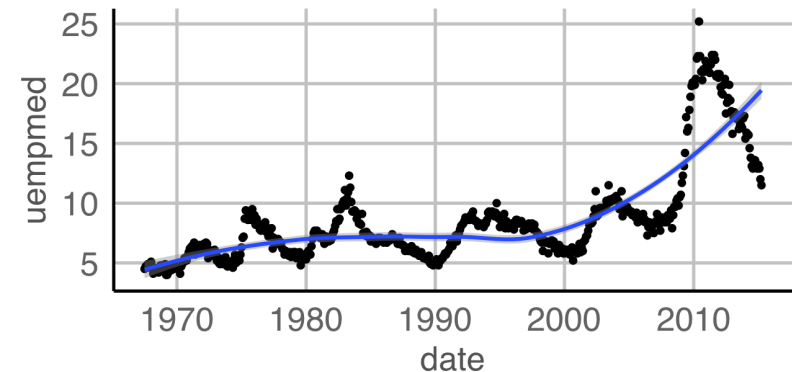
- the default span is 0.75 and
- the default local polynomial degree is 2.

```
fit <- economics %>%  
  mutate(index = 1:n()) %>%  
  loess(uempmed ~ index,  
        data = .,  
        span = 0.75,  
        degree = 2)
```

Showing it on the plot

In `ggplot`, you can add the loess using `geom_smooth` with `method = loess` and method arguments passed as list:

```
ggplot(economics, aes(date, uempmed)) +  
  geom_point() +  
  geom_smooth(method = loess,  
              method.args = list(span = 0.75,  
                                degree = 2))
```



Why non-parametric regression?

- Fitting a line to a scatter plot where noisy data values, sparse data points or weak inter-relationships interfere with your ability to see a line of best fit.
- Linear regression where least squares fitting doesn't create a line of good fit or is too labour intensive to use.
- Data exploration and analysis.
- Recall: In a parametric regression, some type of distribution is assumed in advance; therefore fitted model can lead to fitting a smooth curve that misrepresents the data.
- In those cases, non-parametric regression may be a better choice.
- *Can you think of where it might be useful?*

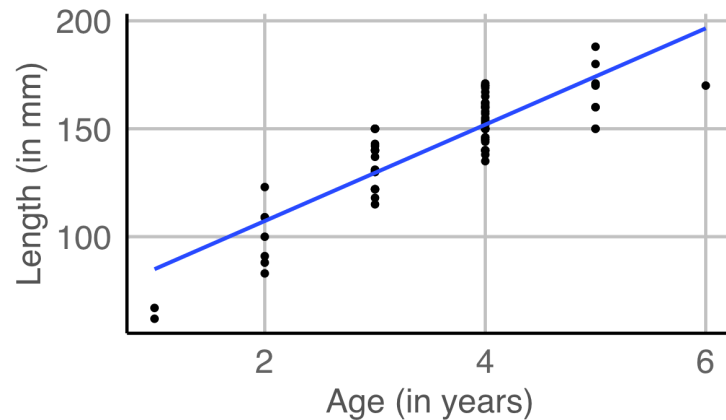
Case study 4 Bluegills Part 1/3

Data were collected on length (in mm) and the age (in years) of 78 bluegills captured from Lake Mary, Minnesota in 1981.

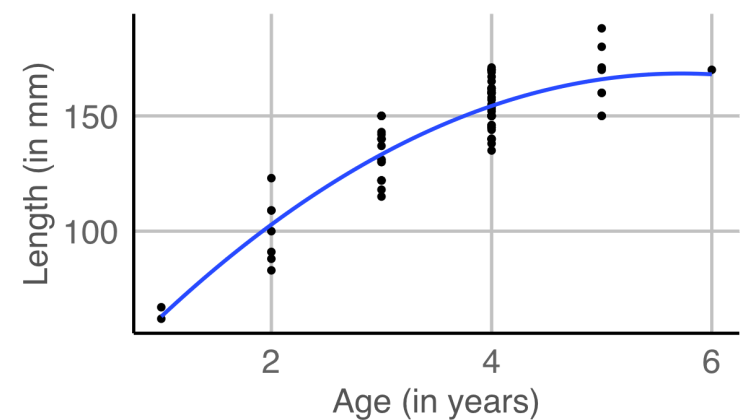
 data R

Which fit looks better?

(A) Linear regression



(B) Quadratic regression



Weisberg (1986) A linear model approach to backcalculation of fish length, *Journal of the American Statistical Association* **81** (196) 922-929

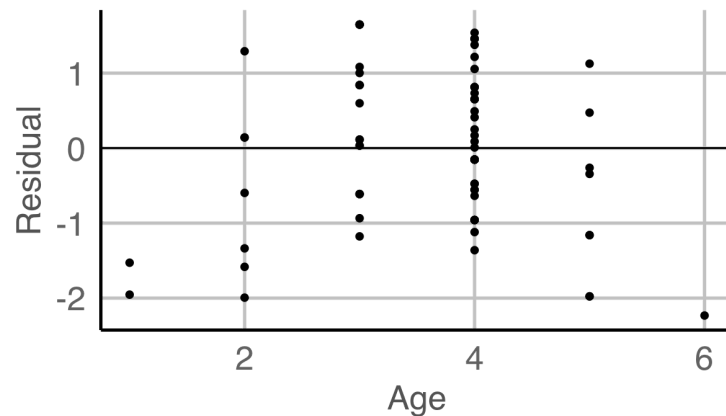
Case study 4 Bluegills Part 2/3

- Let's have a look at the residual plots.
- Do you see any patterns on either residual plot?

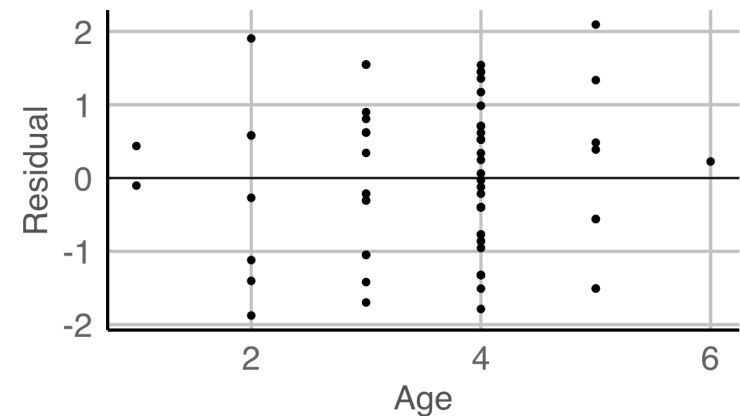


data R

(A) Linear regression



(B) Quadratic regression

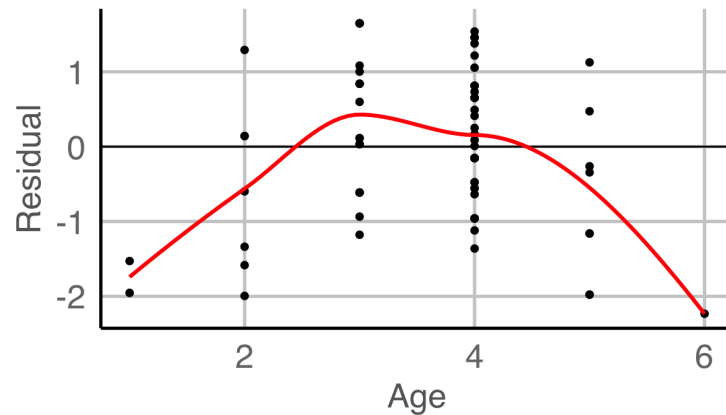


Case study 4 Bluegills Part 3/3

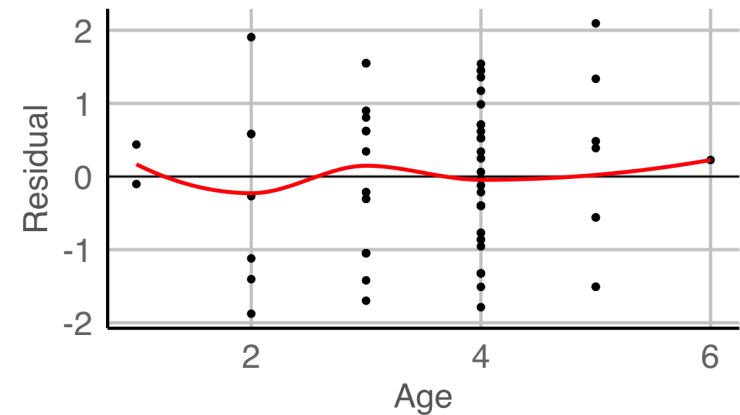
The structure is easily visible with the LOESS curve:

 data R

(A) Linear regression



(B) Quadratic regression



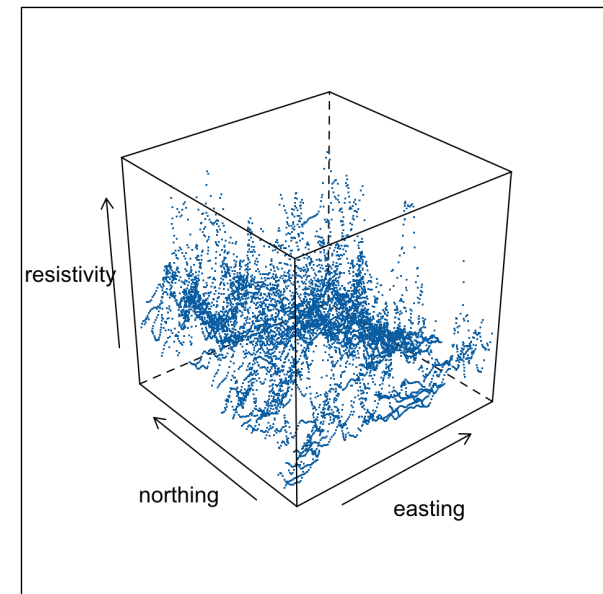
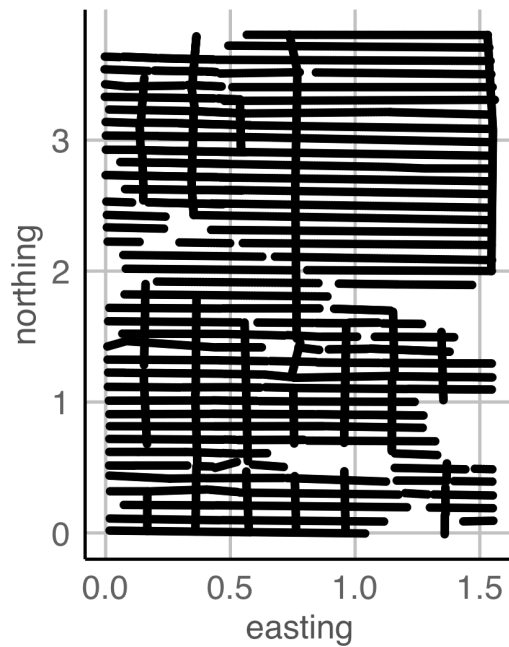
Weisberg (1986) A linear model approach to backcalculation of fish length, *Journal of the American Statistical Association* **81** (196) 922-929

Case study 5 Soil resistivity in a field

This data contains measurement of soil resistivity of an agricultural field.

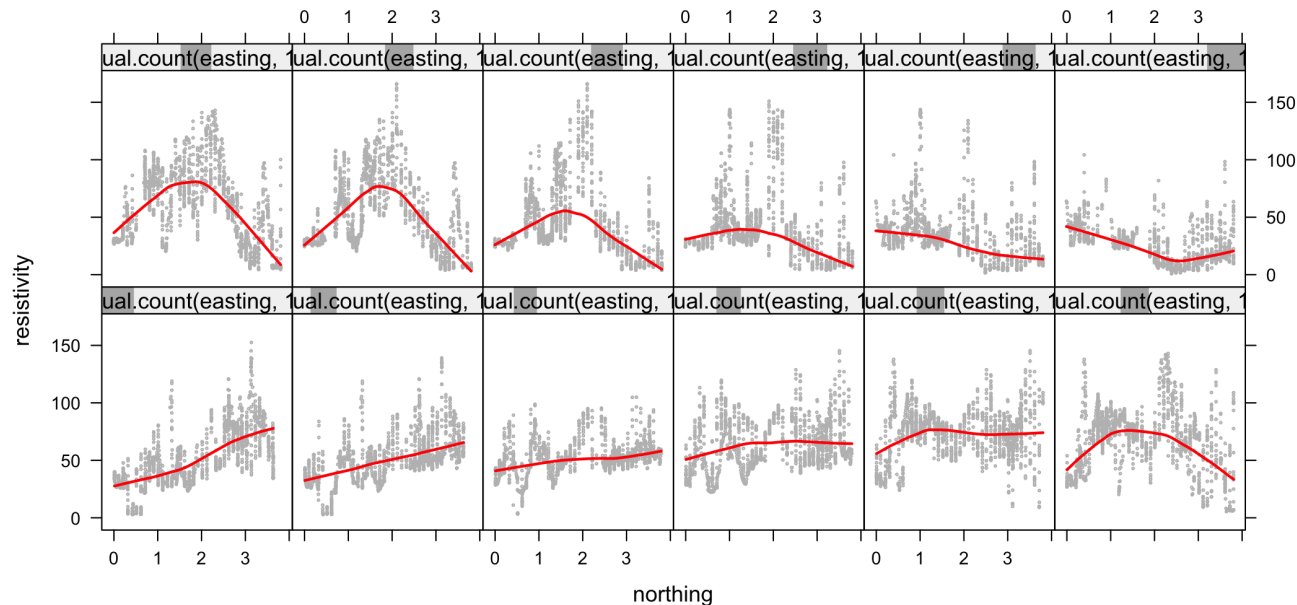


data R



Conditioning plots (Coplots)

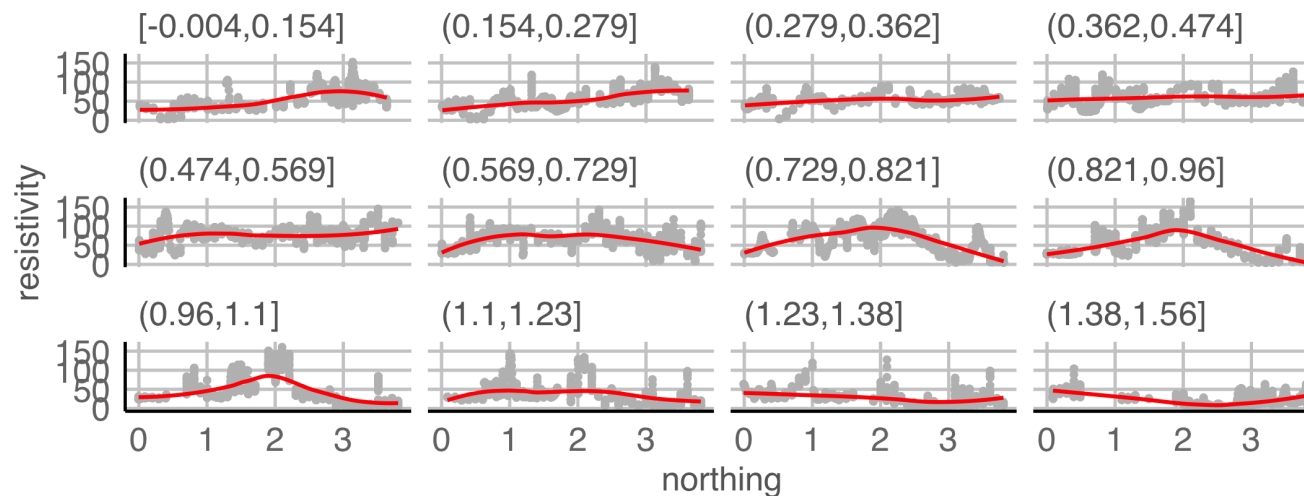
```
library(lattice)
xyplot(resistivity ~ northing | equal.count(easting, 12),
       data = cleveland.soil, cex = 0.2,
       type = c("p", "smooth"), col.line = "red",
       col = "gray", lwd = 2)
```



Coplots via **ggplot2**

- Coplots with **ggplot2** where the panels have overlapping observations is tricky.
- Below creates a plot for non-overlapping intervals of **easting**:

```
ggplot(cleveland.soil, aes(northing, resistivity)) +  
  geom_point(color = "gray") +  
  geom_smooth(method = "loess", color = "red", se = FALSE) +  
  facet_wrap(~ cut_number(easting, 12))
```



Take away messages

- ✈ You can use leverage values and Cook's distance to query possible unusual values in the data
- ✈ Non-parametric regression, such as LOESS, can be useful in data exploration and analysis although parameters must be carefully chosen not to overfit the data
- ✈ Conditioning plots are useful in understanding the relationship between pairs of variables given at particular intervals of other variables

Resources and Acknowledgement

- These slides were originally created by Dr Emi Tanaka, and modified by Dr Michael Lydeamore.
- Cook & Weisberg (1994) "An Introduction to Regression Graphics"
- Data coding using [tidyverse suite of R packages](#)
- Slides constructed with [xaringan](#), [remark.js](#), [knitr](#), and [R Markdown](#).



MONASH
University



This work is licensed under a [Creative Commons Attribution-ShareAlike 4.0 International License](https://creativecommons.org/licenses/by-sa/4.0/).

Lecturer: *Di Cook*

✉ ETC5521.Clayton-x@monash.edu

📅 Week 11 - Session 2

